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## Alcoholic Marriage: Later Start, Sooner End

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### Abstract

**Background**—Although associations between drinking behavior and marital status are well documented, timing of marital transitions as a function of alcohol use or disorder has received limited empirical attention.

**Methods**—We examine the relationship between lifetime history of alcohol dependence (AD) and timing and survival of first marriages in a sample of 3575 female and 1845 male adult Australian twins born mostly between 1940 and 1964. Survival analyses were conducted using Cox proportional hazards regression models.

**Results**—Results indicate moderate delays in marriage associated with AD for both women and men. Among ever married respondents, AD was strongly predictive of early separation, with similar effects observed for women and men. Heritable sources of covariation were also documented. For women, genetic influences shared between early-onset AD and marital timing were found. Genetic influences shared between AD and marital survival were observed for women without regard to onset and for men with later-onset AD.

**Conclusions**—Results confirm the importance of AD as a predictor of both timing and survival of first marriages, with genetic influences contributing to observed associations.

### Keywords

alcohol dependence; marital onset; marital survival; genetic risks

### Introduction

The relationship between marriage and drinking behavior is a topic of much interest in both alcohol and broader health literatures. Higher rates of heavy and problem drinking are consistently observed among never married and divorced individuals, compared to the continuously married (e.g., Chilcoat and Breslau, 1996; Hajema and Knibbe, 1998; Temple et al., 1991; Wilsnack et al., 1984). While questions remain, researchers generally agree that two processes largely account for alcohol-related differences in marital status—nonrandom selection into (and out of) marriage as a function of prior use and a “marriage effect” (Bachman et al., 1997a; Leonard and Rothbard, 1999) related to changing roles and responsibilities associated with marriage.

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Young adults who are heavy drinkers are less likely to marry than either abstainers or light drinkers (Chilcoat and Breslau, 1996; Power and Estaugh, 1990; Wilsnack et al., 1991). At the same time, drinking decreases among individuals who become married (Bachman et al., 1997a, 1997b; Chilcoat and Breslau, 1996; Curran et al., 1998; Hajema and Knibbe, 1998; Hanna et al., 1993; Harford et al., 1994; Power and Estaugh, 1990; Temple et al., 1991) or engaged to be married (see Bachman et al., 1997a; Miller-Tutzauer et al., 1991), with the reverse observed of never married individuals, whose drinking often continues or even increases. Given that marital discord is high among heavy and problem users (for review, see Marshal, 2003), it is not surprising that when they do marry, heavy and problem users are more likely to divorce (Caces et al., 1999; Caetano and Clark, 1998; Collins et al., 2007). Becoming unmarried is also associated with increased drinking (Bachman et al., 1997b; Chilcoat and Breslau, 1996; Hajema and Knibbe, 1998; Magura and Shapiro, 1988; Power and Estaugh, 1990; Power et al., 1999; Temple et al., 1991) and this is especially true of the recently separated (but see Wilsnack et al., 1991).

To date, relatively little is known regarding the impact of alcohol involvement on timing of marital transitions across development (see Leonard and Eiden, 2007). Several studies document associations between adolescent substance use and early marriage or cohabitation (Chassin et al., 1992; Martino et al., 2004; Newcomb and Bentler, 1985, 1988; Power and Estaugh, 1990; Yamaguchi and Kandel, 1985), but not all are consistent with respect to drinking as opposed to smoking or illicit drug use (e.g., Martino et al., 2004). Few studies report on marital onset past the early- to mid-20s, and still fewer use methods taking into account that individuals who have not yet married (or divorced), may do so in the future.

Among notable exceptions, Fu and Goldman (1996) examined marital onset as a function of a range of health behaviors in the National Longitudinal Survey of Youth (NLSY), a prospective cohort of youth initially assessed at ages 14–21. Results of survival analyses indicated significant delays in first marriage associated with heavy drinking observed over a 12-year period through the early 30s. Compared to abstainers and light drinkers, heavy drinking women experienced approximately 40% lower rates of marriage. A similar pattern was found for heavy drinking men, with marriage rates ranging between 34–37% lower, compared to controls. In a subsequent report on the same sample, Fu and Goldman (2000) found few significant effects of alcohol consumption on timing of divorce examined across a 14-year period.

Forthofer et al. (1996) also used survival analysis to examine associations between history of psychiatric disorder and marital onset. In a sample of 15–54 year-olds drawn from the National Comorbidity Study (NCS; Kessler et al., 1994), alcohol dependence (AD) was unrelated to early marriage (defined as marriage prior to age 19), “on-time” marriage (between ages 19–24), or later marriage (after age 24). However, a greater percentage of individuals with substance use disorders, including AD, were considered “off time” in either direction, relative to those with affective or conduct disorders. Given that later marriage was defined as marriage occurring after age 24, delays associated with a later cut-off, for example, age 30 or 35, remain unknown. Forthofer et al. (1996) did not examine risks for divorce, but a second report on the same sample found AD to be a moderate predictor of marital duration, with much stronger effects for women than men (Kessler et al., 1998).

In the present study, we examine the relationship between lifetime history of AD and both marital onset and survival in an adult cohort of Australian twins spanning a wide range of ages, from 28–92, born mostly between 1940 and 1964. We include important control variables to ensure specificity of observed AD effects, ranging from sociodemographic characteristics, history of psychopathology, to family and childhood risks. Among sociodemographic characteristics are educational attainment, a strong predictor of marital

delay (Moore & Waite, 1981), and age at childbearing. Age at childbearing is included given well-documented associations with early marriage, but also potential protective effects of children against separation (Cherlin, 1977; Waite & Lillard, 1991). We further extend previous work by examining heritable sources of observed AD-marital transition covariation. Genetic variation in AD is considerable for both men and women (Heath et al., 1997; Kendler et al., 1994; McGue, 1994), with genetic variation in marital status also reported (Johnson et al., 2004; McGue and Lykken, 1992; Jockin and McGue, 1996; Trumbetta et al., 2007). Although heritable covariation has been suggested (Osler et al., 2008; Prescott and Kendler, 2001), whether there is genetic overlap between AD and timing of marital transitions is currently unknown.

## Materials and Methods

### Participants

Respondents were twins born between 1893 and 1964, with approximately 80% of the sample born between 1940–1964, drawn from a broadly representative Australian volunteer twin panel maintained by the Australian National Health and Medical Research Council (NHMRC). Twins are of primarily European decent reflecting the predominantly Caucasian Australian population from which the cohort was ascertained. While well-educated individuals are over-represented, any impact of differential participation rates as a function of educational attainment appears minimal (Heath et al., 1997; Heath et al., 2001), as does potential bias with respect to drinking behavior (Jardine and Martin, 1984).

Twins were recruited as adults and first surveyed by mailed questionnaire over the period 1980–1981 (Heath et al., 1997). During 1988–89, pairs for whom both twins returned the 1980–81 questionnaire were asked to complete a follow-up questionnaire replicating most of the original items, with additional assessment of female reproductive health. In 1992, all twins from the 1988–89 sample (along with a small sample of twins from an earlier laboratory study; Martin et al., 1985a; Martin et al., 1985b) were asked to complete a structured diagnostic interview. Twins were selected for analysis if they had data on lifetime DSM-III-R alcohol dependence (AD) symptoms assessed during the 1992 interview and variables used to code marital onset, which were included in the 1988–89 questionnaire, but omitted from interview. Of 5879 individual twins completing 1988–89 questionnaire and interview, 5420 (92%) had data on both AD and marital onset, including 3575 female and 1845 male respondents. This includes data from 2338 monozygotic (MZ) and dizygotic (DZ) twin pairs ( $n_{\text{female MZ}} = 828$ ,  $n_{\text{male MZ}} = 345$ ,  $n_{\text{female DZ}} = 461$ ,  $n_{\text{male DZ}} = 184$ ,  $n_{\text{opposite-sex DZ}} = 520$ ) and 744 non-pair twins. Age at interview ranged from 28 to 92 [ $M$  ( $SD$ ) = 44.41 (12.25)].

### Measures

Twins completed a 20-page self-report questionnaire assessing health behavior, including alcohol use and smoking histories, and a telephone adaptation of the Semi-Structured Assessment of the Genetics of Alcoholism (SSAGA; Bucholz et al., 1994; Hesselbrock et al., 1999). The SSAGA was developed for the Collaborative Study on the Genetics of Alcoholism (COGA) to assess physical, psychological, and social manifestations of alcohol abuse or dependence and related psychiatric disorders in adults and is based on previously validated research interviews (e.g., DIS, CIDI, HELPER, SAM, SADS, and SCID). Trained interviewers, who were supervised by a project coordinator and clinical psychologist, administered all interviews. Interviews were tape-recorded and a random sampling of tapes was reviewed for quality control and coding inconsistencies. Informed consent was obtained from all participants prior to their participation using procedures approved by the

institutional review boards at both Washington University School of Medicine and Queensland Institute of Medical Research.

Measures were drawn from interview assessments unless otherwise indicated. A summary of individual measures follows, with descriptive statistics presented in Table 1 by cohort and respondent sex. Where available, age of onset is also shown.

**Marital Onset**—Marital onset was computed from date of first marriage or when the respondent began living in a de facto (common law) partnership. Respondents were asked to provide dates of up to three marriages or partnerships in the 1988–89 questionnaire only. Marital history was assessed only briefly during interview, with respondents who were married or currently living with a partner as though married asked to report years of marriage to current partner, regardless of any prior separation. Years of marriage were not assessed of separated or divorced and widowed respondents not currently remarried and/or living with a partner.

**Marital Survival**—Age at separation from first marital partner was coded from year first marriage or de facto partnership ended for respondents reporting history of separation or divorce. As with marital onset, dates of separation from first and subsequent marriages or partnerships were assessed in the 1988–89 questionnaire only.

**Alcohol Use Disorder**—Lifetime history of AD was coded according to DSM-III-R criteria, with age of onset directly assessed (age at first 12-month clustering of three or more AD symptoms). The SSAGA administered predates publication of DSM-IV criteria. Although an approximate algorithm for DSM-IV AD without clustering criteria has been developed (see Heath et al., 2001), data on age of onset of DSM-IV AD are incomplete as a consequence of DSM-III-R assessment. DSM-III-R was instead analyzed given the importance for the present analysis of ensuring that onset of AD precede first marriage and subsequent separation.

**Control Variables**—Important correlates of AD were selected as control variables based on review of the literature and availability in the current sample, including, for comparability, control variables examined in previous analyses of AD and reproductive timing (Waldron et al., 2008).

**Sociodemographic characteristics:** In addition to birth year, coded in 5-year intervals, educational attainment and religiosity were included among sociodemographic control variables, all of which were drawn from the 1988–89 questionnaire. Dummy variables for leaving high school prior to grade 12 and completing high-school (e.g., earning the equivalent of a high school diploma) were computed with any tertiary education comprising the reference group. Respondents reporting any versus no religion were coded positive for religious affiliation. Weekly or more church attendance was coded from respondent report of frequency of church attendance using a scale ranging from 1 = “more than once a week” to 5 = “rarely”, with 2 = “once a week.”

**Reproductive history:** Age at first childbirth was computed by subtracting respondent’s date of birth from the date of birth of his or her first-born biological child. In adjusted models predicting marital survival, age at childbirth was dummy-coded from time-varying person-year data as age at first, second, and third childbirth, with respondents reporting no children comprising the reference group. Respondents were asked to provide dates of birth of up to ten biological children in the 1988–89 questionnaire only.

**Smoking history:** History of regular smoking was assessed only in the 1988–89 questionnaire. Respondents who self-described as “ex-” or “current smoker”(s) were coded positive for regular smoking. Age of onset of regular smoking was not assessed.

**History of psychopathology:** Childhood conduct disorder (CD), lifetime history of major depressive disorder (MDD), and suicide attempt and panic attack histories were included among psychiatric control variables, each assessed as part of the SSAGA. Relaxed criteria were used to diagnose CD, defined as three or more DSM-IV CD symptoms each with onset before age 18. Likewise, the broader construct of panic attack, rather than panic disorder, was used. DSM-IV criteria were applied to DSM-III-R assessment of MDD and panic attack. History of attempted suicide was also assessed.

**Family and childhood risks:** Maternal and paternal alcohol problems were coded from respondent report of biological parent history of alcohol-related problems with health, family, job, or police, or other problems. Childhood sexual abuse (CSA) was defined as unwanted or forced sexual activity prior to age 18.

**Zygosity—**Zygosity was diagnosed based on twins’ responses to standard questions regarding similarity and the degree to which others confused them (Nichols and Bilbro, 1966). Diagnoses derived from extensive blood sampling have been shown to demonstrate 95% agreement with similar questionnaire-based zygosity determination (Martin and Martin, 1975; Ooki et al., 1990).

## Analytic Strategy

The present study analyzes data from individual twins to examine phenotypic associations between marital timing and AD and among ever married individuals, time to separation and AD. Pair-level data were analyzed to examine potential heritable sources of covariation. Because not all respondents had aged through periods of highest likelihood of marriage or separation, time-to-event data were analyzed using survival analysis. Analyses were conducted separately for women and men given well-documented sex differences in the prevalence of AD and sex differences in marital onset, with women marrying at earlier ages on average than men (ABS, 2009; U.S. Census, 2010).

In preliminary analyses, cumulative failure curves were estimated using the Kaplan-Meier survivor function (Kaplan and Meier, 1958), with log-rank tests to identify significant differences in equality of survivor functions by sex. We also tested for zygosity differences that might identify limitations to the generalizability of twin data. Because monozygotic (MZ) twinning occurs at random, MZ twins will represent a near-random sample of genotypes in the general population. In contrast, dizygotic (DZ) twinning shows weak associations with factors including maternal age and socioeconomic status (Bulmer, 1970).

To examine marital onset as a function of AD history, without and with adjustment for control variables, Cox proportional hazards regression (Cox, 1972) was used. AD was modeled as a time-varying predictor to ensure onset before or at the same time as first marriage. This was achieved using person-year data, with each line of data representing a single year of life for every respondent. For respondents reporting history of AD, AD was coded as absent in each year up to the age they first met criteria and present from that year onward. Among ever married individuals, AD was modeled as a time-varying predictor of age at separation from first marital partner, with respondents entering the analysis (or risk set) at marital onset and widow/ers right-censored upon spousal death. Age at childbirth was also examined as a time-varying predictor in adjusted models, with respondents reporting biological children coded as having none in each year up to childbirth and having one or

more from year of childbirth onward. In analyses of marital onset, age at first childbirth was included as control for premarital childbearing. Age at first, second and third childbirth were included to control for number of biological children born prior to or in the same year as marital separation. The Efron approximation (Efron, 1977) was used for survival ties.

Given the wide age range of respondents, we also tested for age by AD interactions. For these analyses, birth year was re-coded in 5-year increments over the range of 1945–1959 and birth years prior to 1945 were collapsed into a single dummy, with the most recent cohort born between 1960–1964 comprising the reference group. Post-hoc tests were conducted to examine whether any of the birth-year dummy variables could be equated followed by tests for interactions with AD.

Analyses of genetic influences contributing to observed associations between AD and marital timing and survival were conducted using data from complete MZ and DZ twin pairs. The following variables were modeled as predictors in genetically-informed Cox analyses: respondent AD, zygosity, cotwin AD, the interaction between zygosity and cotwin AD (a test of genetic effects), cotwin sex, and the interaction of zygosity, cotwin AD, and cotwin sex (a test of sex differences in genetic effects). Respondent history of AD was modeled as a time-varying predictor, with cotwin AD included as time-invariant. To examine differential effects of early versus later onset, cotwin AD was also dummy-coded as occurring before or on versus after the median age of onset for women or men (ages 24 and 21, respectively), with non-alcoholic cotwins comprising the reference group. To examine the overall genetic contribution to covariation, without regard to univariate heritability, respondent AD was excluded from genetically-informed models of marital timing and survival in separate analyses.

Results from Cox models are presented as Hazard Ratios (HRs). In Cox regression, the dependent variable is called the hazard, which describes event occurrence over time or the rate of event occurrence. A hazard rate is a conditional instantaneous event rate, calculated as a function of time. For dichotomous predictors, HRs represent a ratio of hazard rates for two groups; that is, the ratio of rates at which events are occurring in one group relative to a reference group. HRs greater than 1.0 indicate a higher rate of event occurrence, thus in the context of the present study, earlier marriage compared to the reference group. HRs less than 1.0 indicate a lower rate, suggesting later or delayed marriage compared to the reference group. HRs equal to 1.0 indicate no difference in rates or risk related to onset of marriage. The same is true for marital separation, i.e., HRs greater and less than 1.0 indicate higher and lower rates of separation, respectively.

All survival analyses were conducted using STATA version 8.2 (StataCorp, 2005), with the Huber-White robust variance estimator used to compute standard errors and confidence intervals adjusted for non-independence (i.e., the correlated nature) of data on twin pairs. An important assumption of Cox regression is that of proportional hazards. Violations of proportional hazards, such as might be the case if AD effects on marital onset or survival differ for earlier versus later age periods, were examined using the Grambsch and Therneau test of Schoenfeld residuals (Grambsch and Therneau, 1994), which involves testing for a non-zero slope in a regression of scaled Schoenfeld residuals on functions of time for each predictor variable. Following Cleves, Gould and Gutierrez (2004), interactions between age or risk period and predictor variables were modeled to correct observed proportional hazards violations. Where there was no violation in proportional hazards, risk periods were collapsed.

## Results

### Preliminary Analyses

As shown in Table 1, more women than men reported having ever married. Results of log-rank tests indicate significant sex differences in the cumulative probability of first marriage, with women marrying earlier on average than men. Among ever married individuals, an approximately equal proportion of women and men separated from their first marital partner. Sex differences in the cumulative probability of separation were nonsignificant. Consistent with previous reports, rates of AD were higher for men than women, with onset of AD occurring earlier on average for men. For control variables, sex differences were significant at  $p < 0.01$  unless otherwise noted.

Tests for zygosity differences in marital onset, marital survival, and AD provide support for generalizability of twin data to singleton populations. Nonsignificant differences by zygosity were observed in the cumulative probability of first marriage (women:  $\chi^2_1 = 2.38$ ,  $p = .12$ ; men:  $\chi^2_1 = .12$ ,  $p = .73$ ) and separation among ever-married individuals (women:  $\chi^2_1 = 1.17$ ,  $p = .28$ ; men:  $\chi^2_1 = 2.98$ ,  $p = .08$ ). Zygosity was also unrelated to AD (women:  $\chi^2_1 = 1.77$ ,  $p = .18$ ; men:  $\chi^2_1 = 3.40$ ,  $p = .07$ ).

### Cox Analyses of Marital Timing and Survival

HRs and 95% CIs from models unadjusted and adjusted for sociodemographic, reproductive, psychiatric, and family and childhood risks are presented in Tables 2 and 3 for marital onset and survival, respectively. Not shown are results from models testing for birth-year by AD interactions, which were nonsignificant in all models examined.

**Marital Onset**—For both women and men, there was a moderate association between AD and delayed marital onset. For women, onset of AD was associated with 23% decreased likelihood of first marriage, with no significant age interaction. In adjusted models, 31% decreased likelihood of first marriage was observed, but through age 29. For men, onset of AD was associated with marital delay only after age 29, with 36% decreased likelihood of first marriage. However, in adjusted models, across age, effects of AD for men were nonsignificant.

Several control variables were also robust predictors of marital onset. Through age 19, having a child was associated with nearly 16 times the likelihood of first marriage for women and over 40 times for men. After age 19, risks reduce but remain high for both women and men. Dropping out of high school predicted early marriage for women only. Effects were especially pronounced through age 19, with dropping out was associated with 9 times the likelihood of first marriage. Completing high school without continued schooling predicted early marriage for both women and men, although the relationship was considerably weaker for men. For women, weekly church attendance predicted a slight delay in marital onset, with little to no effect observed for men. Among psychiatric risks, MDD was associated with 16% increased likelihood of first marriage for women through age 24. For men, past suicide attempt was associated with 88% decreased likelihood of first marriage through age 24. Of significant family and childhood risks, maternal alcoholism predicted early marriage for men, with 62% increased likelihood. Regular smoking was associated with earlier marriage for men only.

**Marital Survival**—For both women and men, there was a strong association between AD and early separation. Likelihood of marital separation among alcoholic women was over two times that of non-alcoholics, with no significant age interaction, but this effect disappeared in adjusted models. For men, onset of AD was also associated with over two times the

likelihood of separation. Significant albeit somewhat reduced effects were observed in adjusted models, with 84% increased likelihood of marital separation.

Among control variables, age at first childbirth was strongly predictive of marital survival, but for women only. Through age 24, having a single child was associated with over three times the likelihood of separation, compared to women without children. Having a single child was protective against separation from age 25 onwards as was having two or three or more children. While educational attainment was unrelated to marital survival for women, high school drop-out was associated with a nearly two times increased likelihood of separation for men through age 34. Religious affiliation and weekly church attendance strongly predicted delayed separation for both women and men. History of MDD was associated with over two times the likelihood of separation for women after age 24 and for men, across age, over three times. Past suicide attempt was also predictive of early separation for both women and men—for women, a 78% increased likelihood was observed and for men, nearly four times. Regular smoking was associated with earlier separation for women and non-significant but trend-level effects were observed for men.

### Analyses of Genetic Covariation

Hazard Ratios and 95% CIs from Cox proportional hazards models of genetic covariation between AD and marital onset and survival are shown in Tables 4 and 5, respectively. Not shown are results from models that exclude respondent AD, which were nearly identical to those reported below.

**Marital Onset**—For both women and men, genetic effects were nonsignificant in models of marital timing where cotwin AD was examined without regard to onset (see Model 1). Neither the interaction of zygosity and cotwin AD nor the interaction of zygosity, cotwin AD and cotwin sex were significant, suggesting the absence of omnibus genetic and sex-dependent genetic effects, respectively. In models with cotwin AD coded as occurring on or before versus after the median age of onset (Model 2), genetic effects were observed, but for women only. For women, the interaction of zygosity and cotwin AD with onset before age 25 predicted delayed marital onset, with the interaction of zygosity and cotwin AD and cotwin sex nonsignificant.

**Marital Survival**—In models of genetic covariation between marital survival and AD, sex-dependent genetic effects were observed for both women and men. For women, the interaction of zygosity, cotwin AD and cotwin sex predicted early separation (Model 1), with female like-sex pairs showing earlier separation than opposite-sex pairs. For men, the interaction of zygosity and cotwin AD with onset after age 21 predicted early separation (Model 2). The addition of cotwin sex to the interaction indicates significant sex differences in the direction of men from like-sex pairs showing dramatically earlier separation those from opposite-sex pairs.

### Discussion

Although associations between drinking behavior and marital status are well documented, as are changes in drinking as a function of entry into (and out of) marriage, timing of marital transitions as a function of alcohol use or disorder has received only limited empirical attention. In the present study, we examined the relationship between lifetime history of alcohol dependence (AD) and onset of first marriage, and among ever married individuals, survival of first marriage. We used survival analysis to gauge overall likelihood of marriage and separation, as well as timing of these events, including control variables with well-documented association with AD. Capitalizing on the availability of genetically-informed

twin data, we also examined potential heritable sources of covariation contributing to observed associations.

Consistent with Fu and Goldman (1996), a pattern of delayed marital onset was observed among alcoholic compared to non-alcoholic individuals. For women, AD was associated with reduced likelihood of marriage regardless of adjustment for important sociodemographic, reproductive, psychiatric, and family and childhood risks; a significant age-interaction was found in adjusted models, with delays observed through age 29. For men, marital delays associated with AD were observed in unadjusted models, but only after age 29. Regarding marital survival, consistent with Kessler et al. (1998), among ever married women and men, a pattern of early separation was found without age-interaction. In unadjusted models, likelihood of separation was nearly double for alcoholic compared to non-alcoholic individuals. However, for ever married women, the association between AD and marital survival reduced to non-significance in adjusted models.

Together, results suggest AD may work to delay marriage for women during earlier risk periods, i.e., their 20s, when most women in this cohort married. Although men on average married only a year or two later, marital delay associated with AD appears more pronounced for men starting around age 30. Such differences may be related to the relative rarity of AD among women in this cohort, with AD perhaps representing a more severe phenotype than for men, and this would be especially true of early onset AD, i.e., AD preceding marital onset in the 20s. In contrast, effects of AD on marital survival do not appear to differ by risk period, although reduced statistical power associated with low prevalence outcomes, such as separation (with 12–14% prevalence in this cohort), may be a factor.

As noted, AD remained predictive of marital onset for women and marital survival for men in the presence of highly correlated sociodemographic, reproductive, psychiatric, and family and childhood risks. This includes birth year, a surprising finding given major changes in Australian divorce law easing a number of restrictions, including financial barriers, for obtaining divorce (Donaldson, 2004). In 1959, a system of fault-based divorce allowed dissolution of marriage upon obtaining proof of marital fault, with a single no-fault option available to those who were separated for more than five years. Australia moved to an entirely no-fault system in 1975, providing that the marriage was “broken down irretrievably,” with spouses having lived separately for at least twelve months.

Two control variables, in particular, were strongly predictive of marital timing—educational attainment and age at first childbirth. Leaving school prior to grade 12 was associated with early marriage for women, as was completing high school without continued schooling, together consistent with the larger literature on educational correlates of early versus later marriage (Moore and Waite, 1981). For men, completing high school was associated with early marriage, with no effect of leaving school prior to grade 12. Educational attainment was also significant in analyses of marital survival, but for men only—leaving school prior to grade 12 was associated with early separation through age 34, with no effect of high school completion school without continued schooling.

Childbirth was a powerful predictor of marital onset for both women and men, and this was especially true prior to age 20. While effects were strongest during the teen years, childbirth continued to predict early marriage from age 20 onwards. Interestingly, for women, having a single child was associated with increased likelihood of separation through age 24, but protective against separation from age 25 onwards. Having two or three or more children was equally protective, consistent with research suggesting reduced risks of marital separation where children are involved (Cherlin, 1977; Waite and Lillard, 1991). For men, having children was unrelated to marital survival.

In a previous report, we documented a strong relationship between AD and delayed reproduction among women, with a negligible effect observed for men (Waldron et al., 2008). Although timing of marriage and reproduction are highly correlated, AD was associated with delayed marriage for both sexes. It is possible that onset of marriage and reproduction are more closely linked for alcoholic women than men. Alcoholic women may be less likely to marry, and thus less likely to reproduce, particularly in the present cohort, where early marriage is relatively normative. Of course, reproductive impairments from prolonged or excessive drinking compound such effects (Emanuele et al., 2002), perhaps explaining why AD-related reproductive delays are much stronger than those observed for marriage, at least among women.

Other control variables were also predictive of marital timing, including comorbid psychopathology. For women, major depressive disorder (MDD) was associated with early marriage through age 24 and separation from age 24 onwards. MDD was unrelated to marital onset for men, but strongly predictive of early separation. Attempted suicide was associated with marital delay for men through age 24, and early separation for both women and men. Owing to limited assessment, we examined lifetime history of both MDD and attempted suicide and thus are unable to establish direction of effects with regard to marital timing, but also AD. Depression and AD are highly comorbid and disentangling cause from consequence, especially in light of alcohol-involved marriages, merits further examination.

Although other yet unmeasured covariates may be important, there are several possible explanations of observed marital delays, assuming AD contributes uniquely to timing of first marriages. Alcoholic women and men may be less attractive to potential mates, making for less suitable partners; they may also be less likely to commit to conventional relationships, particularly during younger years. There is also the possibility that alcoholic individuals choose to remain single for longer. By remaining single, regardless of reason, risks for AD may increase as a consequence of continued alcohol use in part because of continued participation in alcohol-oriented singles events and/or stress associated with not marrying, including both social and financial.

Partner effects are also likely. In addition to assortative mating (Agrawal et al., 2006; Grant et al., 2007; Labouvie, 1996), there is growing evidence of spousal concordance for drinking over the marriage course (Demers et al., 1999; Leonard and Eiden, 1999; Leonard and Mudar, 2004), with higher rates of marital dissatisfaction observed among spouses discordant for heavy or problem drinking (Homish and Leonard, 2007; Leadly et al., 2000; Roberts and Leonard, 1998), and not surprisingly, increased risks of divorce (Ostermann et al., 2005). Unfortunately, drinking histories of first spouses were not available for individuals reporting repartnering following separation, divorce or widowhood. However, in follow-up analyses of ever-married individuals, inclusion of current or former spouse history of heavy drinking did little to reduce the magnitude of AD associations with either onset or survival of first marriages (data not shown).

Regardless of mechanism, genetic risks appear to play an important role, especially for women. From analyses of twin pairs, we observed heritable sources of covariation between AD and both marital timing and survival. Results indicate genetic influences contribute to associations between AD and marital delay for women, particularly during early risk periods, consistent with phenotypic observations. For men, heritable covariation between AD and marital onset was nonsignificant, although confidence intervals were wide so as not to preclude the possibility. For both women and men, genetic influences contributed to associations between AD and marital survival. For women, heritable covariation with early separation was found regardless of risk period and for men, during later risk periods.

What underlies genetic results? Among likely candidates are a range of inherited traits, including comorbid psychopathology. Shared genes account at least partially for associations between personality and both AD (Slutske et al., 2002) and marital status (Jockin et al., 1996; Johnson et al., 2004). Genetic covariation between AD and depression has also been reported (Fu et al., 2002; Prescott et al., 2000). Although a genetically-informed analysis of depression and marital status has yet to be published, relationships between depression and both never marrying and marital instability have long been documented (Horwitz, White & Howell-White, 1996; Waite, 1995).

Although ours is the first genetically-informed analysis of AD effects on timing of marital transitions, we acknowledge several limitations, including reliance on retrospective reports for AD especially. We also know little regarding the role of remission. In the present study, history of AD was analyzed conditioned on three or more AD symptoms experienced within a 12-month period either prior to or at the same time as marital onset or separation. Although effects of AD on marital onset for men reduced to nonsignificance in the presence of control variables, we cannot say that active AD has no impact on likelihood of marriage. A similar argument can be made for AD effects for women in models of marital survival. It is certainly likely that individuals who are actively dependent, compared to those in remission, are at much higher risk for both marital delay and subsequent separation.

There are also limits to the generalizability of findings. First, our sample was drawn from an adult cohort of Australian twins, born between 1893 and 1964, many of whom were coming of age during a socially conservative era, when marriage was more common, especially at younger ages. Cohabitation without marriage (or the common-law equivalent) is increasingly common in Australia today (ABS, 2009). Although some researchers (e.g., Horwitz and White, 1998) have found that both women and men who cohabit report more alcohol problems, compared to married women and men, it is largely unknown whether differences between marriage forms are important. Furthermore, prevalence of marital separation or divorce is very low in this cohort owing in part to restrictive divorce laws prior to 1975. Whether AD contributes to delay in onset of subsequent marriages also remains unknown. Thus, the present findings require replication in a younger sample, where marriage may be less common and differences associated with nonmarital cohabitation or remarriage can be explored. Although data on AD from a younger cohort of twins born 1964–1971 is available, as of last assessment (see Heath et al., 2001; Knopik et al., 2006), the majority remain either unmarried or non-cohabiting (Waldron et al., 2009). Follow-up of this cohort will be important for purposes of replication, but also generalization of findings.

In addition, it is possible that observed patterns differ for populations other than our predominantly Caucasian sample. This is especially true given differences by race in risk for substance use or disorder, including AD, with prevalence often lower among African Americans (Grant, 1997; Grant and Dawson, 1997; Hasin et al., 2007). Likelihood of marriage and subsequent separation differs by race as well (U.S Census Bureau, 2010). Cross-national differences are also possible, with onset of marriage and marriage-like relationships varying widely even among developed countries (Corijn and Klijzing, 2001; Fussell and Gauthier, 2004). The same may be true of AD, although international comparisons of AD prevalence are difficult because of diagnostic differences and differences in whether prevalence for lifetime versus past year prevalence is reported (WHO, 2004).

Despite limitations, the present study contributes to current understanding of the relationship between marriage and alcohol involvement by extending previous research to examine AD associations with both onset and survival of first marriage, including tests of heritable covariation. Results confirm the importance of AD as a moderate to strong predictor of both

marital delay and early separation, with genetic influences contributing to observed relationships. To better understand the nature of AD differences in timing of marital transitions, additional research on correlated genetic as well as environmental risks is required. Studies including time-varying comorbid psychopathology as well as personality may be particularly informative, as may studies of other substance use or disorder. Although regular smoking was only weakly related to marital timing, the role of comorbid illicit drug use remains unknown.

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**Table 1**

Descriptive statistics, by respondent gender

	<b>Women (n = 3575)</b>	<b>Men (n = 1845)</b>
Ever married	3160 (88)	1532 (83)
Age at first marriage, <u>M</u> ( <u>SD</u> )	23.21 (3.94)	25.53 (4.23)
Ever separated/divorced from first marital partner, <u>n</u> (%)	437 (14)	187 (12) <sup>NS</sup>
Age at first separation, <u>M</u> ( <u>SD</u> )	32.55 (8.49)	33.66 (9.02) <sup>NS</sup>
Time (years) to separation, <u>M</u> ( <u>SD</u> )	10.63 (8.07)	9.15 (8.08) <sup>NS</sup>
AD, <u>n</u> (%)	204 (6)	430 (23)
Age of onset, <u>M</u> ( <u>SD</u> )	27.52 (8.82)	24.35 (7.72)
<b>Control Variables</b>		
Biological children, <u>n</u> (%)	2729 (77)	1243 (67)
Age at first childbirth, <u>M</u> ( <u>SD</u> )	25.40 (4.04)	28.00 (4.26)
<b>Educational attainment</b>		
Early school leaver, <u>n</u> (%)	1221 (34)	292 (16)
High school, no tertiary education, <u>n</u> (%)	1311 (37)	699 (38) <sup>NS</sup>
<b>Religiosity</b>		
Religious affiliation, <u>n</u> (%)	2932 (87)	1376 (80)
Weekly church attendance, <u>n</u> (%)	849 (25)	372 (22)
Regular smoking, <u>n</u> (%)	1473 (41)	909 (49)
CD, <u>n</u> (%)	81 (2)	305 (17)
MDD, <u>n</u> (%)	778 (22)	288 (16)
Suicide attempt, <u>n</u> (%)	109 (3)	37 (2)
Panic attack, <u>n</u> (%)	145 (4)	41 (2) <sup>a</sup>
Maternal alcoholism, <u>n</u> (%)	132 (4)	54 (3) <sup>NS</sup>
Paternal alcoholism, <u>n</u> (%)	602 (17)	299 (17) <sup>NS</sup>
CSA, <u>n</u> (%)	198 (6)	47 (3)

**Note:** AD = alcohol dependence; CD = childhood conduct disorder; MDD = major depressive disorder; CSA = childhood sexual abuse;

<sup>NS</sup> gender difference not significant at  $p = .05$ ;

<sup>a</sup>  $p = 0.02$ .

Table 2

Hazard Ratios (and 95% CIs) from Cox proportional hazards models of marital onset

Predictor (Risk Period)*	Women		Men	
	Unadjusted	Adjusted	Unadjusted	Adjusted
AD (<30)	0.77 (0.63 – 0.96)	0.69 (0.52 – 0.92)	0.96 (0.81 – 1.14)	0.86 (0.71 – 1.05)
AD (≥30)		1.74 (0.93 – 3.28)	0.64 (0.43 – 0.95)	
<u>Control Variables</u>				
Birth year (<20)		1.13 (1.07 – 1.18)		1.01 (0.98 – 1.04)
Birth year (20–25)		1.02 (1.00 – 1.04)		
Birth year (≥25)		0.93 (0.91 – 0.96)		0.92 (0.88 – 0.94)
Age at first childbirth (<20)		15.95 (10.92 – 23.29)		41.71 (8.83 – 196.93)
Age at first childbirth (20–24)		4.12 (3.21 – 5.28)		6.78 (4.29 – 10.71)
Age at first childbirth (≥25)		3.03 (2.11 – 4.35)		4.46 (2.95 – 6.73)
HS drop-out (<20)		9.02 (6.00 – 13.57)		1.08 (0.09 – 1.30)
HS drop-out (20–24)		1.74 (1.51 – 2.01)		
HS drop-out (25–29)		1.05 (0.83 – 1.33)		
HS drop-out (≥30)		3.29 (1.65 – 6.57)		
HS degree (<20)		3.65 (2.46 – 5.44)		1.17 (1.03 – 1.34)
HS degree (20–24)		1.27 (1.11 – 1.44)		
HS degree (25–35)		1.03 (0.83 – 1.27)		
HS degree (≥35)		1.85 (1.30 – 2.66)		
Religious affiliation		1.10 (0.96 – 1.26)		1.04 (0.88 – 1.23)
Weekly church attendance		0.86 (0.78 – 0.95)		0.97 (0.84 – 1.12)
Regular smoking		0.94 (0.86 – 1.02)		1.27 (1.12 – 1.43)
CD		0.98 (0.66 – 1.45)		1.18 (0.99 – 1.39)
MDD (<25)		1.16 (1.04 – 1.30)		1.01 (0.86 – 1.19)
MDD (≥25)		0.92 (0.76 – 1.13)		
Suicide attempt (<25)		1.04 (0.78 – 1.40)		0.82 (0.34 – 1.96)
Suicide attempt (≥25)		1.61 (0.95 – 2.71)		0.22 (0.08 – 0.60)
Panic attack		1.09 (0.88 – 1.35)		1.05 (0.67 – 1.65)
Maternal alcoholism		1.02 (0.81 – 1.30)		1.62 (1.22 – 2.15)
Paternal alcoholism		0.96 (0.85 – 1.09)		1.00 (0.85 – 1.18)
CSA		1.00 (0.83 – 1.22)		0.76 (0.53 – 1.09)

Note:

\* risk period in years of age; AD = alcohol dependence; HS = high school; CD = childhood conduct disorder; MDD = major depressive disorder; CSA = childhood sexual abuse.

**Table 3**

Hazard Ratios (and 95% CIs) from Cox proportional hazards models of marital survival

Predictor (Risk Period)*	Women		Men	
	Unadjusted	Adjusted	Unadjusted	Adjusted
AD	<b>2.42 (1.54 – 3.82)</b>	0.95 (0.55 – 1.64)	<b>2.45 (1.76 – 3.40)</b>	<b>1.84 (1.22 – 2.77)</b>
<u>Control Variables</u>				
Birth year		<b>1.16 (1.10 – 1.22)</b>		<b>1.10 (1.02 – 1.20)</b>
1 child (<25)		<b>3.10 (1.79 – 5.37)</b>		†
1 child (≥25)		<b>0.58 (0.39 – 0.88)</b>		1.08 (0.67 – 1.75) ‡
2 children		<b>0.61 (0.44 – 0.85)</b>		1.03 (0.63 – 1.69)
3+ children		<b>0.57 (0.40 – 0.80)</b>		0.64 (0.35 – 1.18)
HS drop-out (<35)		†		<b>1.82 (1.05 – 3.16)</b>
HS drop-out (≥35)		1.09 (0.82 – 1.46) ‡		0.92 (0.42 – 2.00)
HS degree		1.12 (0.85 – 1.46)		1.26 (0.87 – 1.83)
Religious affiliation		<b>0.66 (0.50 – 0.89)</b>		<b>0.49 (0.33 – 0.73)</b>
Weekly church attendance (<25)		†		2.01 (0.74 – 5.43)
Weekly church attendance (≥25)		<b>0.52 (0.39 – 0.71)</b> ‡		<b>0.51 (0.27 – 0.95)</b>
Regular smoking		<b>1.49 (1.20 – 1.85)</b>		1.39 (0.96 – 2.01)
CD (<25)		†		2.41 (0.98 – 5.89)
CD (≥25)		1.19 (0.65 – 2.20) ‡		0.67 (0.39 – 1.15)
MDD (<25)		1.02 (0.57 – 1.85)		†
MDD (≥25)		<b>2.29 (1.81 – 2.91)</b>		<b>3.13 (2.20 – 4.46)</b> ‡
Suicide attempt		<b>1.78 (1.12 – 2.84)</b>		<b>3.78 (1.81 – 7.90)</b>
Panic attack (<35)		†		2.49 (0.84 – 7.37)
Panic attack (≥35)		0.78 (0.48 – 1.27) ‡		~0
Maternal alcoholism		1.22 (0.72 – 2.04)		0.71 (0.26 – 1.93)
Paternal alcoholism		1.03 (0.79 – 1.35)		1.37 (0.89 – 2.13)
CSA		1.21 (0.79 – 1.87)		0.77 (0.25 – 2.34)

Note:

\* risk period in years of age; AD = alcohol dependence; HS = high school; CD = childhood conduct disorder; MDD = major depressive disorder; CSA = childhood sexual abuse.

**Table 4**

Hazard Ratios (and 95% CIs) from Cox proportional hazards models of genetic covariation between AD and marital onset

	Women	Men
<b><u>Predictors (Risk Period)*</u></b>		
Model 1		
Respondent AD (<30)	T	0.99 (0.82 – 1.19)
Respondent AD (≥30)	0.80 (0.64 – 1.00)	<b>0.64 (0.42 – 0.97)</b>
	⊥	
Zygoty	1.07 (0.87 – 1.32)	1.25 (0.88 – 1.78)
Cotwin AD	1.16 (0.66 – 2.02)	1.55 (0.81 – 2.95)
Zygoty * cotwin AD	0.41 (0.07 – 2.51)	0.18 (0.02 – 2.06)
Cotwin sex	1.12 (0.98 – 1.28)	0.93 (0.78 – 1.10)
Zygoty * cotwin AD * cotwin sex	1.36 (0.70 – 2.64)	1.95 (0.72 – 5.32)
Model 2		
Respondent AD (<30)	T	0.99 (0.82 – 1.19)
Respondent AD (≥30)	0.83 (0.67 – 1.03)	<b>0.64 (0.42 – 0.97)</b>
	⊥	
Zygoty	1.08 (0.88 – 1.33)	1.18 (0.83 – 1.68)
Cotwin early AD	2.62 (1.00 – 6.86)	1.45 (0.58 – 3.64)
Cotwin later AD	0.85 (0.41 – 1.75)	1.05 (0.45 – 2.47)
Zygoty * cotwin early AD	<b>0.04 (0.002 – 0.82)</b>	0.35 (0.01 – 10.54)
Zygoty * cotwin later AD	0.82 (0.07 – 9.06)	0.62 (0.02 – 17.66)
Cotwin sex	1.11 (0.97 – 1.27)	0.95 (0.80 – 1.13)
Zygoty * cotwin early AD * cotwin sex	2.25 (0.81 – 6.24)	1.48 (0.37 – 5.86)
Zygoty * cotwin later AD * cotwin sex	1.29 (0.52 – 3.19)	1.25 (0.31 – 5.00)

**Note:**

\* risk period in years of age; AD = alcohol dependence.

**Table 5**

Hazard Ratios (and 95% CIs) from Cox proportional hazards models of genetic covariation between AD and marital survival

	Women	Men
<b><u>Predictors (Risk Period)*</u></b>		
Model 1		
Respondent AD	<b>2.22 (1.36 – 3.63)</b>	<b>2.48 (1.73 – 3.54)</b>
Zygoty	0.89 (0.54 – 1.45)	1.04 (0.35 – 3.08)
Cotwin AD	0.57 (0.13 – 2.43)	1.21 (0.24 – 6.14)
Zygoty * cotwin AD	2.60 (0.02 – 381.36)	14.57 (0.05 – 4625.88)
Cotwin sex (<35)	<b>1.56 (1.01 – 2.41)</b>	†
Cotwin sex (≥35)	0.85 (0.51 – 1.43)	0.70 (0.03 – 2.56)
Zygoty * cotwin AD * cotwin sex	<b>1.40 (1.01 – 2.41)</b>	‡
Model 2		
Respondent AD	<b>2.19 (1.33 – 3.61)</b>	<b>2.44 (1.71 – 3.49)</b>
Zygoty	0.90 (0.56 – 1.48)	1.03 (0.36 – 2.98)
Cotwin early AD	0.91 (0.08 – 9.45)	4.35 (0.62 – 30.65)
Cotwin later AD	0.46 (0.07 – 3.11)	0.09 (0.002 – 4.73)
Zygoty * cotwin early AD	0.31 (0.0001 – 836.06)	0.004 (~0 – 76.71)
Zygoty * cotwin later AD	11.58 (0.02 – 7592.85)	<b>250566 (1.11 – &gt; 100)</b>
Cotwin sex (<35)	<b>1.58 (1.02 – 2.42)</b>	†
Cotwin sex (≥35)	0.87 (0.52 – 1.46)	0.71 (0.42 – 1.18)
Zygoty * cotwin early AD * cotwin sex	3.81 (0.19 – 76.03)	‡
Zygoty * cotwin later AD * cotwin sex	0.67 (0.06 – 7.73)	<b>0.01 (0.0001 – 0.55)</b>

**Note:**

\* risk period in years of age; AD = alcohol dependence.